

### **Research Article**

## Statistical Properties and Application of Lindley– Kumaraswamy Distribution

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### **ARTICLE INFO:**

ABSTRACT

Keywords:	In this study, we proposed an extension of the Kumaraswamy					
Distribution,	distribution by adding a parameter to the Kumaraswamy					
Lindley,	distribution using the Lindley-G family of distributions proposed					
Kumaraswamy,	by Cakmakyapan and Ozel (2016). The model parameters were					
Maximum	estimated using Maximum Likelihood Estimation (MLE). The					
Likelihood	performance of the new Lindley-Kumaraswamy distribution was					
G family	compared with other family of Kumaraswamy distribution. The					
2	performance was tested by an application to a real-life data and					
	compared with other family of Kumaraswamy distributions using					
	the Akaike Information Criterion (AIC) and the Bayesian					
	Information Criterion (BIC) which are two types of information					
	criterion. Our new distribution was found to provide a better fit					
	than the existing distributions when modeling positively skewed					
	data.					

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### **INTRODUCTION**

Mathematical flexibility and simplicity of probability distributions are factors to consider in the choice of a statistical model otherwise known as distribution. Experts in probability distribution theory have connected these important features in their choice of distribution for statistical research. This connection had opened another investigation where many researchers put some choosing standard theoretical distributions that best fits real life phenomena. Lindley (1958) was the first to ever conceptualize the idea of the Lindley distribution. A Lindley distribution is one way to portray the duration of a process or device. Since its formulation the distribution has been used in a wide variety of fields, including engineering, medicine and biology. Ghitany et al. (2008) stated its usefulness especially for modeling in mortality studies. Other

$$G(x) = 1 - \left[1 + \frac{\theta x}{\theta + 1}\right] e^{-\theta x}$$
$$g(x) = \frac{\theta^2}{\theta + 1} (1 + x) e^{-\theta x}$$

for interval [0,1], for x > 0,  $\theta > 0$  where  $\theta$  represent scale parameter of the Lindley distribution.

The Kumaraswamy distribution describes a distribution within which outcomes are restricted to a selected interval, the characteristics of probability density function inside this interval depends upon the shape parameters. There is a commonality between Kumaraswamy distribution and another distribution called Beta distribution (Handique and Chakraborty, 2016). But Kumaraswamy is probably easier to use due to its easier analytical expressions both for its probability density function and accumulative distribution function.

The Kumaraswamy distribution was originally conceptualized by Poondi Kumaraswamy in 1980 for random processes that are doubly bounded for applications that are hydrological. The capacity to simulate monotone and non-

(1)

(2)

monotone failure rate functions also lies in this distribution whose commonality in life-time data analysis and consistency appear quite often. Recently, several researches have been carried on different families of Kumaraswamy distributions. For example, Marcelino et al. (2011) introduced and studied the Kumaraswamy generalized gamma distribution which will be used in to model bathtub shaped hazard rate functions. For more applications on Kumaraswamy family of distribution, see for example, Bourguignon et al. (2012), Santana et al. (2012), Mansour et al. (2018), Varghese and Jose (2022), among others.

The Kumaraswamy distribution is denoted by  $K(\alpha, \beta)$  and belongs to the family of continuous probability distributions outlined within the interval [0,1] with a cumulative distribution function *cdf* given by:

$$G(x) = 1 - \left(1 - x^{\alpha}\right)^{\rho}$$

(3)

Also supplied is the matching probability density function pdf.:

$$g(x) = \alpha \beta_x^{\alpha - 1} \left( 1 - x^{\alpha} \right)^{\beta - 1}$$

For interval  $0 \le x \le 1$ , where  $\alpha$  and  $\beta$  indicate the shape characteristics of the Kumaraswamy distribution.

### Motivation

Amongst the flaws in statistical analysis is that most accessible data sets do not adhere appropriately many of the current and popular conventional probability distribution(s) and for that cannot provide an adequate modelling. Our effort in this study is meant to propose a different distribution namely Lindleydistribution Kumaraswamv using Lindley-G family of distributions by adding a parameter in order to obtain a flexible compound distribution which is better than the existing ones. A data set has been used to demonstrate the utility of the new distribution. The distribution has both shape and scale parameters which also make it flexible to model tailed distribution. The study of this distribution is significant, has various  $-\log[1-G(x)]$ 

(4)

uses in areas which involve tailed distribution problems and in engineering, such as in hydrology, and other fields of study. Also, its extension provides a flexible model fitting among its family. **Organization of the Study** 

The work is organized as follows. Section 2 provides materials and methodology of the research. The data analysis is presented in section 3. Finally, section 4 discussion of findings and section 5 brings to conclusion.

### METHODS AND MATERIALS Lindley-Kumaraswamy Distribution

Cakmakyapan and Ozel (2016) proposed a Lindley-Generalized distribution that provides more flexibility in modeling of real-life data. Now, the *cdf* and *pdf* of the Lindley-G family of distributions according to Cakmakyapan and Ozel (2016) were derived from the transformation given by;

$$F(x) = \int_{-\infty}^{-\log[t-O(x)]} \frac{\theta^2}{\theta+1} (1+t) e^{-\theta t} dt$$
(5)

Now, the *cdf* is derived from the transformation by integration by parts as:

$$F(x) = 1 - \left[1 - \frac{\theta}{(\theta+1)} \left[\log(1 - G(x))\right]\right] \left[1 - G(x)\right]^{\theta}$$
(6)

*pdf* is obtained by differentiating the *cdf* as:

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$$f(x) = \frac{\theta^2}{(\theta+1)} g(x) \left[ 1 - \log(1 - G(x)) \right] \left[ 1 - G(x) \right]^{\theta-1}$$

$$\tag{7}$$

respectively. For x > 0,  $\theta > 0$ , G(x) and g(x) represent the cumulative distribution function (cdf) and the probability distribution function (pdf) of the **baseline** or **parent** distribution

respectively and  $\theta$  is the scale parameter of the Lindley-G family of distribution.

Substituting equations (3) and (4) in (6) and (7) respectively, the cumulative distribution function (cdf) and probability distribution function (pdf) of Lindley-Kumaraswamy

distribution

respectively are obtained thus:

$$F(x) = 1 - \left[1 - \frac{\theta}{(\theta+1)} \left[\log(1 - G(x))\right]\right] \left[1 - G(x)\right]^{\theta}$$

$$F(x) = 1 - \left[1 - \frac{\theta}{(\theta+1)} \left[\log\left(1 - \left(1 - (1 - x^{\alpha})^{\beta}\right)\right)\right]\right] \left[1 - \left(1 - (1 - x^{\alpha})^{\beta}\right)\right]^{\theta}$$

$$F(x) = 1 - \left[1 - \frac{\theta}{(\theta+1)} \left[\log\left(1 - x^{\alpha}\right)^{\beta}\right]\right] \left(1 - x^{\alpha}\right)^{\theta\beta}$$
(8)

and

$$f(x) = \frac{\theta^2}{(\theta+1)} g(x) \Big[ 1 - \log(1 - G(x)) \Big] \Big[ 1 - G(x) \Big]^{\theta-1}$$
$$f(x) = \frac{\theta^2}{(\theta+1)} \Big( \alpha \beta_x^{\alpha-1} (1 - x^{\alpha})^{\beta-1} \Big) \Big[ 1 - \log(1 - (1 - (1 - x^{\alpha})^{\beta})) \Big] \Big[ 1 - (1 - (1 - x^{\alpha})^{\beta}) \Big]^{\theta-1}$$

$$f(x) = \frac{\alpha \beta \theta^2}{(\theta+1)} x^{\alpha-1} (1-x^{\alpha})^{\beta(\theta-1)} \left[ 1 - \log(1 - (1 - (1 - x^{\alpha})^{\beta}) \right]$$
(9)

respectively.

For  $0 \le x \le 1$ , where  $\alpha$ ,  $\beta$  are shape parameters and  $\theta$  is a scale parameter. Hence equations (8) and (9) represent the *cdf* and *pdf* of the proposed Lindley-Kumaraswamy Distribution.

**Graphical Presentation of the** *pdf* **and** *cdf* Considering a certain value for parameters  $\alpha$ ,  $\beta$ , and  $\theta$ , we prepared some potential figures for the probability distribution function (pdf) and the cumulative distribution function (cdf) of the LinKumD as demonstrated below:



**Figure 1:** Shape of the probability distribution function (pdf) of the projected distribution for various true parameter values.

**Interpretation:** The figure above shows that the growth of the cdf is in direct proportionality with increase in x and approaches 1 when x gets bigger. This means that the distributions are extremely valuable for data sets with many different shapes.



Figure 2: Shape of the *cdf* of the proposed distribution for different parameters values

**Interpretation** The figure above indicates, as expected, that the cdf rises when *x* rises and approaches 1 as *x* rises. **Estimation of Parameters of the LinKumD** Let  $X_1, X_2, ..., X_n$  be independently and identically distributed random variables

of sample size n of the LinKumD with previously defined unknowns  $\alpha$ ,  $\beta$ , and  $\theta$ . Equation (10) gives the pdf of the LinKumD as:

$$f(x) = \frac{\alpha \beta \theta^2}{(\theta+1)} x^{\alpha-1} (1-x^{\alpha})^{\beta(\theta-1)} \left[ 1 - \log(1-(1-(1-x^{\alpha})^{\beta})^{\beta(\theta-1)}) \right]$$

The probability function is defined as follows:

$$L(X/\alpha,\beta,\theta) = \left(\frac{\alpha\beta\theta^2}{\theta+1}\right)^n \prod_{i=1}^n \left\{ x_i^{\alpha-1} (1-x_i^{\alpha})^{\beta(\theta-1)} \left[ 1 - \log\left(1 - (1-(1-x_i^{\alpha})^{\beta})\right) \right] \right\}$$

$$(10)$$

Let the log-likelihood function,

$$l = \log L(X / \alpha, \beta, \theta)$$

Therefore,

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$$\sum_{i=1}^{n} \log \beta + 2n \log \theta - n \log(\theta + 1) + (\alpha - 1) \sum_{i=1}^{n} \log x_i + (\theta \beta - 1) \sum_{i=1}^{n} \log(1 - x_i^{\alpha}) + \sum_{i=1}^{n} \log \left[ 1 - \log \left( 1 - (1 - (1 - x_i^{\alpha})^{\beta}) \right) \right]$$
(11)

After differentiating *l* partially with respect to  $\beta$  and  $\theta$  respectively gives;

$$\frac{\partial l}{\partial \beta} = \frac{n}{\beta} + \theta \sum_{i=1}^{n} \log\left(1 - x_{i}^{\alpha}\right) + \sum_{i=1}^{n} \left\{ \frac{\left(1 - x_{i}^{\alpha}\right)^{\beta} \ln\left(1 - x_{i}^{\alpha}\right)}{\left[1 - \log\left(1 - x_{i}^{\alpha}\right)^{\beta}\right] \left(1 - x_{i}^{\alpha}\right)^{\beta}} \right\}$$

$$\frac{\partial l}{\partial \theta} = \frac{2n}{\theta} - \frac{n}{\left(\theta + 1\right)} + \beta \sum_{i=1}^{n} \log\left(1 - x^{\alpha}\right)$$
(12)

The solution of the non-linear system of equations;  $\frac{\partial l}{\partial \beta} = 0$  and  $\frac{\partial l}{\partial \theta} = 0$  will provide us with the maximum likelihood parameter estimates of  $\beta$  and  $\theta$ . When data sets are available, the solution can only be computed numerically using proper statistical tools such as R, SAS, Maple, and so on.

### DATA ANALYSIS AND RESULTS

This section compares the performance of the Lindley-Kumaraswamy Distribution (LinKumD) to that of two other models such as Transmuted Kumaraswamy distribution (TKumD) and Kumaraswamy distribution (KumD) using a real dataset. We used **R** statistical package for all the analysis in this research work. The density functions of these distributions are depicted below:

# The Lindley-Kumaraswamy Distribution (LinKumD)

The pdf of the LinKumD obtained from equation (9) is as follows:

:

$$f(x) = \frac{\alpha\beta\theta}{(\theta+1)} x^{\alpha-1} (1 - x^{\alpha})^{\beta(\theta-1)} [1 - \log(1 - (1 - (1 - x^{\alpha})^{\beta}))]$$
(14)

### The Transmuted Kumaraswamy Distribution (TKumD)

Khan *et al.* (2016) defined the *pdf* of the TKumD as follows:

$$f(x;\alpha,\theta,\lambda) = \alpha \theta_{\chi}^{\alpha-1} \left(1 - x^{\alpha}\right)^{\theta-1} \left\{ 1 + \lambda - 2\lambda \left(1 - \left(1 - x^{\alpha}\right)^{\theta}\right) \right\}$$
(15)

### The Kumaraswamy Distribution (KumD)

According to Kumaraswamy (1980), the *pdf* of the Kumaraswamy distribution is given as:

$$f(x;a,b) = ab\chi^{a-1} (1-x^a)^{b-1}$$
(16)

Table 1: Shape measurements of 48 rock samples from a petroleum reservoir

0.0903296	0.189651	0.228595	0.200071	0.280887	0.311646	0.176969
0.464125	0.148622	0.164127	0.231623	0.144810	0.179455	0.276016
0.438712	0.420477	0.183312	0.203654	0.172567	0.113852	0.191802
0.19753	0.163586	0.200744	0.117063	0.162394	0.153481	0.291029
0.133083	0.326635	0.253832	0.262651	0.122417	0.150944	0.204314
0.240077	0.225214	0.154192	0.328641	0.182453	0.167045	0.148141
0.262727	0.161865	0.341273	0.276016	0.230081	0.200447	

This data was extracted from BP research, image analysis by Ronit Katz, University of Oxford and has been used for analysis by Javanshiri *et al.* (2015)

We also provide histogram and density for the data set as shown in the Figure 3 below:



Figure 3: A Histogram and Density Plots for the Rock Sample Dataset

From the descriptive statistics, the histogram and the density plot shows the data set as observed is right-skewed and thus appropriate for distributions that are skewed most especially to the right.

### **Model Selection Criteria**

To test the performance of the above listed distributions, we used the following information criteria: The Akaike Information Criterion (AIC) and the Bayesian Information Criterion (BIC). The statistics are as follows:

$$AIC = -2ll + 2k$$
$$BIC = -2ll + k \log(n)$$

Where ll denotes the log-likelihood function, k represents the number of model parameters and n the sample size. **Bench mark:** The model with the lowest value for these statistics is thought to be the best fit for the data.

Estimates and performance measures of the above distributions are listed in Table 2 for the rock sample data set as follows:

Models	Estimated Parameter Values	Log- likelihood Values	Akaike Informatio n Criterion (AIC)	Consistent Akaike Information Criterion (CAIC)	Bayesian Information Criterion (BIC)	Ranks
LinKum D	α =2.663380	-52.38064	-98.7612	-98.21583	-93.14768	1
	β=5.009926					
	θ=9.004390					
TkumD	^α=0.8324228	4.269415	1453883	15.08428	20.15243	3
	^β=0.5579018					
	^ λ=0.4570070					
KumD	^ α=1.894405 ^ β=9.051699		-81.4609	-81.19424	-77.7185	2
	Estimated	Pdfs for the Datase	et	Estimated Cdfs for the	Dataset	
			2-	in man and a maile		

**Table 2:** The performance of the distribution was calculated using R to obtain the AIC and BIC values of the models based on the data set.



Figure 4: Histogram shows the estimated density plots, as well as the CDFs and PDFs of the fitted distributions to data set.



Figure 5: Probability plots for the fit of the LinKumD, TKumD and KumD based on the data set

### DISCUSSIONS

The MLEs of the parameters for the fitted models for the data set are provided in Table 2. The AIC and BIC statistics for the fitted distributions are also included in the rock samples data set table. Based on the data sets and model selection values, it is obvious that the proposed LinKumD with three parameters outperforms the TKumD and the standard KumD. Furthermore, the estimated pdfs and cdfs in figure 2 and the Q-Q plots in figure 3 clearly support and confirm the results in Tables 2 and 3.

### CONCLUSION

This research work proposed a new Kumaraswamy extension of the Lindleydistribution called Kumaraswamy distribution. We estimated the parameters by using the method of Maximum Likelihood Estimation (MLE). We demonstrated the performance of the new distribution using real life data sets. Results of the model indicated that Lindley-Kumaraswamy distribution outperforms Kumaraswamy Transmuted and Kumaraswamy distributions.

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*Classifications*: 62E10, 62E15, 62E99

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